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# Evaluation of Reaction Parameters of the Phenol Steam Reforming over Ni/Co on ZrO<sub>2</sub> Using the Full Factorial Experimental Design

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Abstract: Full factorial experimental design with 32 runs was used to investigate the significant and interaction variable of the reaction parameters on phenol steam reforming toward hydrogen production. Effects of selected factors on the phenol conversion  $(Y_1)$  and hydrogen yield  $(Y_2)$ were evaluated. These factors were as follows: (A) temperature (500 and 800 °C); (B) feed flow rate (0.16–0.46 mL/min); (C) catalyst weight (0.1–0.3 g); (D) Ni-Co ratio (0–1); and (E) phenol concentration in the feed (2–10 wt %). Ni and Co over ZrO<sub>2</sub> support for catalytic performance of phenol steam reforming (SRP) was prepared by the impregnation method. The result indicated that all the main independent variables had significant influence on the dependent variable of  $Y_1$  and  $Y_2$  with a range of 2.7%–96.8% and 21.4%–72.4%, respectively. Additionally, some interaction variables like AE, BE, *CE*, and *DE* have also influenced the  $Y_1$  and  $Y_2$  responses. This design showed that the best initial conditions that produced maximum  $Y_1$  and  $Y_2$  responses were at 800 °C, 0.16 mL/min feed flow rate, 0.3 g of catalyst, 0 ratio of Ni-Co (Co/ZrO<sub>2</sub>), and 10 wt % of phenol in the feed, where the phenol conversion was predicted to be 94.98% and the hydrogen yield was predicted to be 67.4%. Within the limits the variables were examined, a regression model which well-fitted the experimental data was proposed. The regression model were reduced to simplify and to get the significant regression coefficient with *p*-value less than 0.05.

**Keywords:** reaction parameters; phenol steam reforming; Ni/Co on ZrO<sub>2</sub>; full factorial experimental design

# 1. Introduction

Due to fossil fuels exhaustion, the growth of concern about global warming, and the increase in major air toxics, hydrogen has emerged as one of the best clean energy vectors [1]. H<sub>2</sub> can be produced from maintainable liquid fuels or phenols that consist of around 38 wt % [2] of the unwanted component of pyrolysis oil. However, this component can be transformed into H<sub>2</sub> rich gas, which makes phenol an ideal potential H<sub>2</sub> source. Based on previous research, there are two successive reactions that may take place in the steam reforming of phenol [3]. One of the reactions for H<sub>2</sub> production from phenol is steam reforming of phenol (SRP), which can produce more  $H_2$  yield based on stoichiometry in Equations (1) and (2):

$$C_6H_5OH + 5H_2O \rightarrow 6CO + 8H_2, \ \Delta H^o = 710.91 \text{ kJ/mol}$$
 (1)

$$CO + H_2O \leftrightarrow CO_2 + H_2$$
,  $\Delta H^o = -41.15 \text{ kJ/mol}$  (2)

Ni-based catalysts with a high activity for the phenol steam reforming process have been investigated extensively in recent years [4–6] due to the ability to cleave C–C, O–H, and C–H bonds [7]. In addition, cobalt can assist with C–C bond cleaving at temperatures as low as 400 °C; this shows high production of  $H_2$  and  $CO_2$  [8] due to the highly favourable water gas shift reaction [9–11]. However, there are a few items that can be identified as the catalyst in the reforming of phenol: the conversion of phenol, and the yield of  $H_2$ , CO or CO<sub>2</sub>. Many issues affect this performance, such as the steam to phenol (S/P) ratio, catalyst composition, flow rate of carrier gas, operating temperature, reactant feed rate and space time, preparation method, catalyst particle size, amount of catalyst, and ratio of active metals in multiple active metal catalyst use. Exothermic reactions, high phenol conversion, and the rate of gas production all occur at a high temperature. The same results were found in previous works regarding the temperature effect on phenol conversion. Some researchers have found that the conversion of phenol increased with the increasing of temperature [6,12–14]. Based on the works of Rioche et al. [13] and Shurong et al. [15], the increase in catalyst activity is based on the increasing temperature. They have found that a complete phenol conversion occurs at high temperatures. There has been no research studying the amount of catalyst, the feed reactant flow, or the metal ratio in the steam reforming of phenol. However, some previous works did use other feeds like methanol, acetic acid, glycerol, and alcohol.

Regarding previous works, increasing the space time [16,17] or expanding the contact time by increasing the amount of the catalyst [18,19] can result in an increase in the activity of the catalyst. However, the catalytic performance also depends on the kind of catalyst [17,20,21] and the technique of catalyst preparation [17].

DOE methods are more suitable than one-factor-at-a-time techniques because by studying multiple factors at the same time, a high efficiency can be obtained. These techniques have the following common benefits [22]: (1) more information per experiment than accidental methods; (2) a decrease in the number and fee of experiments; (3) making the calculation of the interaction amongst possible variables within the series studied, leading to better facts about the process; and (4) simplifying the description of the significant operation conditions for the scale-up process. Therefore, these techniques have been considered as powerful tools for process examination [23–27].

A numerical design of experiment (DOE) was applied in order to investigate the interaction amongst factors and the significance of factors on phenol conversion. The DOE capabilities improve the processes by screening the factors to determine which are important for explaining the process variation. After screening the factors, DOE helped to understand how these factors interact during the process. It can then find the factor settings that produced optimal process performance. Thus the DOE method is a great tool for process performance testing [28–30].

The aim of the current study is to investigate the relationship between the five control parameters, significant variables, and the interaction between variables via full factorial design. Therefore, a detailed study of DOE in the catalytic performance of steam reforming of phenol (SRP) for hydrogen production was carried out. The performance of the catalyst was stated according to the phenol conversion and hydrogen yield.

## 2. Experimental

## 2.1. Catalyst Preparation

The impregnation method was applied to produce the Ni or Co over  $ZrO_2$  catalyst. The advantage of the impregnation method is that is produces a high concentration of active materials on the catalyst surface [31]. The procedure of this method was reported by Athanasios et al. [32]. The techniques used to make the catalyst were as follows: firstly, 9 g of Zirconium oxide ( $ZrO_2$ ) was dispersed in 250 mL of deionized water and stirred for 2 h at 90 °C. While stirring the slurry, 4.94 g of active metal (nickel nitrate hexahydrate (Ni(NO<sub>3</sub>)<sub>2</sub>·6H<sub>2</sub>O) or cobalt nitrate hexahydrate (Co(NO<sub>3</sub>)<sub>2</sub>·6H<sub>2</sub>O) was added to the slurry. The mixture was held until it became an extremely viscous paste. The paste was dried at 110 °C for 24 h and the catalyst was calcined at 750 °C for about three hours. Then, the catalyst was pelletized and sieved on two layers of 35 and 34 mesh to acquire particles sizes between 1.0 mm and 1.4 mm.

## 2.2. Catalyst Characterization

The crystalline structure of the reduced catalyst was determined by X-ray diffraction (High-Resolution X-ray Diffractometer brand Bruker D8 Advance, BRUKER AXS INC., Fitchburg, WI, USA) using a Cu K $\alpha$  radiation at 40 kV and 30 mA. Diffraction angles were measured in steps of 0.02°, at 1 s/step in the range of 20 to 80 (2 $\theta$ ). The size of the metal crystallites was calculated from line broadening with the help of the Scherrer equation [33]. The total surface area of catalysts was measured using the multipoint BET-N<sub>2</sub> on a surface area analyzer (Micromeritics, Gemini 2360, Norcross, GA, USA). Prior to the analysis, the samples were degassed at 200 °C for 2 h to remove unwanted components on the surface before outgassing the samples.

A chemisorption analyser (Micromeritics, Chemisorb 2720, Norcross, GA, USA), was used to evaluate the reducibility of the active metals (nickel and cobalt) on the support by applying a temperature-programmed reduction of hydrogen (TPR-H<sub>2</sub>). Before the reduction, 25 mg of the catalyst was treated at 300 °C under high purity 99.99% helium with the flow rate of 20 mL/min for an hour to remove moisture and other gas impurities. The TPR-H<sub>2</sub> profile was obtained by ramping up the temperature at 10 °C/min, 20 mL/min of 10% (vol.) H<sub>2</sub>/Ar between 50 °C and 800 °C.

Thermogravimetric analysis (TGA) profiles were carried out in a Perkin Elmer TGA instrument (Waltham, MA, USA) operated under nitrogen flowing in the heating rate of 10  $^{\circ}$ C/min. Exothermic weight loss was observed at the temperature range between 30 and 950  $^{\circ}$ C.

## 2.3. Catalyst Performance Test

A 0.2 g weight of catalyst was employed to investigate the performance of the catalyst. The catalyst bed was located inside a 6-mm ID quartz tubular reactor. A coaxially centred thermocouple was incorporated into the reactor. The nitrogen stream passed through the catalyst at 300 °C and subsequently reduced in situ with a 30 mL/min stream of pure hydrogen for 1 hour at 600 °C. The mixture of phenol and water was fed into the pre-heater by means of an HPLC pump (Bio-RadTM, Series 1350, Berkeley, CA, USA) before mixing with carrier N<sub>2</sub>. The samples in reduced forms were tested in the steam reforming of phenol (H<sub>2</sub>O/phenol(mol/mol) = 9:1) at an atmospheric pressure and a temperature range of 500 to 800 °C, maintaining the reaction for 6 h in order to check if there was any deactivation of the catalysts. The reaction products were analysed on-line by GC with TCD (Agilent 6890N, Agilent Technologies, Santa Clara, CA, USA ) equipped with Carboxen Plot 1010 capillary column (Fused Silica, 30 m × 0.53 mm, Sigma-Aldrich Corporation, St. Louis, MO, USA) connected in series, using argon as the carrier gas. The 95% level of confidence interval was chosen for the uncertainty in figure values for gas production and phenol conversion. The flow diagram of this work was the same as the study done by Abdullah et al. [34] and is shown in Figure 1. As a result of the

products variation, the catalyst performance was characterized by two parameters; phenol conversion percentage and product yield. They were calculated according to Equations (3) and (4):

Phenol conversion (%) = 
$$\frac{[Phenol]_{in} - [Phenol]_{out}}{[Phenol]_{in}} \times 100$$
 (3)

$$H_2 \text{ yield } (\%) = \frac{\text{moles of } H_2 \text{ obtained}}{\text{moles of } H_2 \text{ stochiometric potential}} \times 100.$$
(4)



Figure 1. Schematic flow diagram of experimental setup.

#### 2.4. Statistical Design of Experiment

A full experiment was carried out in order to study the relationship between the control variables, significant variables, and the interactions between the test parameters. A 2<sup>5</sup>-factorial design was applied in the present investigation. This required a selection of appropriate responses, factors, and levels. Phenol conversion and hydrogen yield were the responses selected. The factors and the experimental levels for each factor were selected based on literature values, available resources, and results from preliminary experiments. The selection of the factors was based on the operating conditions that had a significant influence on the SRP, according to previous experiments (with different feeds) [12–14,16–18]. The considered factors chosen were temperature, A, Feed flow rate, B, Catalyst amount, C, the catalyst ratio of Ni to Co over ZrO<sub>2</sub>, D, and phenol concentration in the feed, E. The factors and their levels were assigned in Table 1. A Minitab statistical software (<sup>®</sup>16.2.4, Minitab, Inc., Pennsylvania State University, University Park, PA, USA, 2013) was used to generate the design matrix and to conduct the data analysis. Also, the interaction among the factors was determined using the same software. The design matrix of un-coded values for the factors and the response for all the experimental runs are given in Table 2. Five factors were analyzed at two levels, using a full factorial  $2^5$  resulting in 32 runs, in a random order to minimize the effect of uncontrolled variables on responses  $(Y_1 \text{ and } Y_2).$ 

Since one of the aims of this study was to establish a relationship between the five control variables and the responses of the system (phenol conversion and hydrogen yield) a linear regression model was fitted to predict the responses as a function of independent variables and their interactions. In general, the response for the linear regression model is described in Equation (5):

$$Y_{i} = \beta_{0} + \beta_{1}A + \beta_{2}B + \beta_{3}C + \beta_{4}D + \beta_{5}E + \beta_{6}AB + \beta_{7}AC + \beta_{8}AD + \beta_{9}AE + \beta_{10}BC + \beta_{11}BD + \beta_{12}BE + \beta_{13}CD + \beta_{14}CE + \beta_{15}DE + \beta_{16}ABC + \beta_{17}ABD + \beta_{18}ABE + \beta_{19}ACD + \beta_{20}ACE + \beta_{21}ADE + \beta_{22}BCD + \beta_{23}BCE + \beta_{24}BDE + \beta_{25}CDE + \beta_{26}ABCD + \beta_{26}ABCE + \beta_{28}ABDE + \beta_{29}ACDE + \beta_{30}BCDE + \beta_{31}ABCDE + \varepsilon_{4}$$
(5)

where  $Y_i$  is the predicted response,  $\beta_0$  is the offset term,  $\beta_n$  is the regression coefficient that corresponds to the main effects and interaction effects, and  $\varepsilon$  is the error term. The level of significance of each coefficient was considered by comparing the smaller *p*-value (<0.05) or greater *t*-value. The *t*-test was used to determine the significance of the regression coefficients of the parameters. The *p*-values were used as tools to check the significance of each variable and the interaction among the variables. As the first step, a significance test was performed on all the possible terms in models for phenol conversion ( $Y_1$ ) and hydrogen yield ( $Y_2$ ). The results are illustrated in Pareto charts and Normal plots. Then a reduced model was obtained by removing the insignificant terms from the first model with a new reduced regression equation. The  $R^2$  value was used to determine the fitting quality of the observed results to the proposed model. The regression analysis, linear regression plots (observed versus predicted value), and Pareto chart of the dependent variables, main effect and interaction plots, surface and contour plots, and response optimizer plots were plotted.

Table 1. Selected factors and levels for the experimental study.

Factors	Symbol	Lev	vels
Temperature (°C)	Α	500	800
Feed flow rate (mL/min)	В	0.16	0.46
Catalyst amount (g)	С	0.1	0.3
Ni-Co ratio	D	0	1
Phenol Concentration (wt %)	Ε	2	10

 Table 2. Experiment matrix for the 2<sup>5</sup> factorial designs and experimental results.

Runs	A (°C)	B (mL/min)	<i>C</i> (g)	D	E (wt %)	Phenol Conversion	$H_2$	CO	CO <sub>2</sub>
1	500	0.46	0.1	0	10	21.6	31.6	10.4	58
2	500	0.46	0.1	1	10	21.0	31	10	59
3	800	0.46	0.3	1	10	53.4	41.1	12.8	46.1
4	800	0.46	0.3	1	2	13.9	27.6	8.9	63.5
5	800	0.16	0.1	1	10	64.0	54.2	15.3	30.5
6	800	0.16	0.1	1	2	17.3	29.5	9.6	60.9
7	500	0.16	0.1	1	2	9.4	25.1	6.8	68.1
8	500	0.16	0.1	0	2	13.2	27.5	8.6	63.9
9	800	0.46	0.1	1	10	42.1	34.6	12.5	52.9
10	800	0.46	0.1	0	2	11.6	26.9	7.8	65.3
11	800	0.16	0.1	0	10	81.9	62.2	15.9	21.9
12	500	0.16	0.1	1	10	37.6	33.9	11.9	54.2
13	800	0.46	0.1	0	10	57.9	51.2	14.8	34
14	800	0.16	0.3	0	2	23.2	32.5	11.2	56.3
15	500	0.46	0.3	1	10	22.0	32.3	10.6	57.1
16	500	0.46	0.3	0	10	62.6	53.9	14.9	31.2
17	500	0.16	0.3	1	10	44.0	37.5	12.6	49.9
18	500	0.16	0.3	0	2	11.6	27.1	8.5	64.4
19	800	0.16	0.1	0	2	19.7	30.7	9.7	59.6
20	500	0.46	0.3	0	2	15.6	28.1	9	62.9
21	500	0.46	0.1	1	2	5.0	22.3	5.9	71.8
22	800	0.16	0.3	1	2	16.3	28.4	9.2	62.4

Runs	A (°C)	B (mL/min)	C (g)	D	E (wt %)	Phenol Conversion	$H_2$	CO	CO <sub>2</sub>
23	800	0.16	0.3	0	10	96.8	72.4	16.1	11.5
24	800	0.46	0.1	1	2	9.7	25.2	7.4	67.4
25	500	0.16	0.1	0	10	57.4	45.9	13.5	40.6
26	500	0.16	0.3	0	10	58.1	51.5	14.8	33.7
27	500	0.46	0.1	0	2	2.7	21.4	5.8	72.8
28	500	0.46	0.3	1	2	5.7	22.4	6.2	71.4
29	800	0.16	0.3	1	10	76.3	55.7	15.7	28.6
30	800	0.46	0.3	0	10	77.9	56.8	15.8	27.4
31	500	0.16	0.3	1	2	10.6	26.7	7.7	65.6
32	800	0.46	0.3	0	2	19.5	30.1	9.7	60.2

Table 2. Cont.

## 3. Results and Discussion

#### 3.1. Characterisation of the Fresh Catalysts

XRD analysis was used to identify the crystal phases of the catalysts. The XRD patterns of the Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts after reduction at 600 °C are displayed in Figure 2. The XRD spectrum was characterized by several intense peaks between the diffraction angles of 20° and 80°. The XRD pattern of the Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts contains Bragg peaks at around  $2\theta = 44.52^{\circ}$ , which corresponds to metallic Ni. This can be attributed to the presence of the cubic [JCPDS 45-1027] structure of Ni and is in agreement with previous research [6,35–37]. The diffraction peaks at  $2\theta = 44.37^{\circ}$  correspond to metallic Co [JCPDS 01-1254], similar to reports from previous works [38,39]. It can be inferred that nickel oxide (NiO) and cobalt oxide (CoO) were reduced at 600 °C to metallic Ni [40,41] and metallic Co [42], respectively. The relative intensity of XRD peaks of the supported catalysts depends on the nature of the metal oxide precursors formed during the calcination step [43]. The XRD patterns of the ZrO<sub>2</sub> support resulted in a number of peaks located at 20 of 24.47°, 28.19°, 31.48°, 34.45°, 40.73°, 50.61°, 55.62°, 58.3°, 60.02°, 62.83°, 69.67°, and 71.29° corresponding to the (111), (200), (210), (211), (300), (320), (400), (411), (331), (332), and (422) crystal planes, which can be attributed to the monoclinic ZrO<sub>2</sub> (m-ZrO<sub>2</sub>) and tetragonal ZrO<sub>2</sub> (t-ZrO<sub>2</sub>) [JCPDS 83-0943] structure of the ZrO<sub>2</sub>. This conforms with previous research [44–47].



**Figure 2.** XRD pattern for Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts after reduction at 600 °C; (•) peaks correspond to metallic Ni, ( $\blacksquare$ ) peaks correspond to metallic Co, ( $\square$ ) peaks correspond to monoclinic phase of ZrO<sub>2</sub> (m-ZrO<sub>2</sub>) and ( $\bigcirc$ ) peaks correspond to tetragonal phase of ZrO<sub>2</sub> (t-ZrO<sub>2</sub>).

To have better insights into the nature of the cobalt and nickel phases, TPR-H<sub>2</sub> of the calcined catalysts was carried out. This helped to explore the level of reduction of the Co and Ni oxides' species. Figure 3 shows that only single reduction zones were observed. This corresponds to the area below 400 °C (~280–450 °C), corresponding to the reduction of NiO [48,49] and CoO [50–52] particles with weak metal–support interaction. Thus, the reduction of the NiO and CoO to metallic Ni and Co agrees with the XRD analysis. In particular, the reductions were peaks shifted to higher temperatures as the Co/ZrO<sub>2</sub> catalyst. This suggests that higher temperatures are needed in order to reduce the cobalt oxides (>452 °C) completely, as compared to nickel oxides [53]. Therefore, these catalysts were reduced at 600 °C.



Figure 3. TPR-H<sub>2</sub> profiles for the Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts.

## 3.2. Factors Screening in a Full 2<sup>5</sup> Factorial Design for the Initial Regression Model

Table 2 gives the experimental values of the two main responses in each experiment. The 2-level factorial was employed to improve the correlation between the five process variables, temperature, feed flow rate, amount of catalyst and the catalyst of Ni to Co ratio, and phenol concentration in the feed (coded as *A*, *B*, *C*, *D*, and *E*, respectively) to responses of phenol conversion ( $Y_1$ ) and hydrogen yield ( $Y_2$ ). Thirty-two sets of experiments were performed to explore the effects of the reaction parameters on the phenol conversion and the generation of hydrogen in the phenol steam reforming process. It was found that the conversion of the phenol varied from 2.7% to 96.8% and H<sub>2</sub> yield was in the range of 21.4%–72.4%. The maximum catalyst activity for  $Y_1$  of 96.8 and  $Y_2$  of 72.4% was found to be at 800 °C, 0.16 mL/min feed flow rate, 0.3 g of catalyst, 0 ratio of Ni-Co (Co/ZrO<sub>2</sub>), and 10 wt % of phenol in the feed (run number 32).

To identify whether the results are "real" or "chance," a Normal plot (Figure 4) was used to display the estimates of the main effects among the factors as well as their interaction. To obtain this Normal plot, effects were ranked in ascending order and plotted against their corresponding cumulative probability [54]. All the effects that tended to fall along a straight line in the plot (distributed with mean zero and marked in black) do not demonstrate any significant effect on the responses. In comparison, the significant effects were located far from the straight line and had non-zero means. They appear in red. The significant effects get more distanced from the straight line as they become larger. From Figure 4, it can be seen that all of the main parameters (A, B, C, D, and E) and their interactions (AE, BE, CE, and DE) are found to be "real" parameters for  $Y_1$  and  $Y_2$ . The red marks, which lie to the left of the fitted line, had a negative effect on the responses [55].



**Figure 4.** Normal plot for (**a**)  $Y_1$  and (**b**)  $Y_2$ .

The evaluation of the effect of each variable on the  $Y_1$  and  $Y_2$  may also be verified from the Pareto chart. A Pareto chart (Figure 5) visually represents the absolute values of the effects of main factors and the effects of *n*-way interaction of factors. The chart includes a vertical reference line at the critical *t*-value for an alpha of 0.05 (magnitude for a 95% confidence level) to indicate that the factors that extend past this line are potentially important. For a 95% confidence level and 32 degrees of freedom (DF), *t*-value is equal to 4.89 for  $Y_1$  and 3.04 for  $Y_2$ . It can again be seen from the Pareto chart that all effects are statistically significant. The significance of the *n*-way interactions could also be compared from the Pareto charts. These results were in agreement with those obtained from the analysis of the Normal plot. These plots clearly indicate that the significance of the variables and their *n*-way interaction towards  $Y_1$  and  $Y_2$  decreased in this order: E > A > B > D > AE > C > DE > BE > CE and E > A > B > D > DE > AE > C > BE.



**Figure 5.** Pareto chart for (a)  $Y_1$  and (b)  $Y_2$ .

Based on the experimental data, regression models were fitted for  $Y_1$  and  $Y_2$ , as shown in Equations (6) and (7), respectively. The adequacy of the initial model was tested a via parity plot for observed versus predicted values, as demonstrated in Figure 6. As seen in Figure 6, the high value of the correlation coefficient ( $R^2 = 1$ ) demonstrates good correlation between the observed and the predicted responses by the initial models. The analysis of the regression coefficient for initial models  $Y_1$  and  $Y_2$  is shown in Tables 3 and 4. In such analysis, a modal term is considered significant if its *p*-value

is less than 0.05 (bold). It was found that among the variables and their interaction, *A*, *B*, *C*, *D*, *E*, *AE*, *BE*, *CE*, and *DE* for *Y*<sub>1</sub> and *A*, *B*, *C*, *D*, *E*, *AE*, *BE*, and *DE* for *Y*<sub>2</sub> response were statistically significant.

$$Y_{1} = 8.74 - 0.002A - 39.7B - 74.48C - 19.41D + 7.49E + 0.02AB + 0.084AC + 0.03AD + 0.001AE + 247.6BC + 65.85BD - 30.93BE + 133.44CD - 30.32CE - 6.78DE - 0.2ABC - 0.09ABD + 0.03ABE - 0.21ACD + 0.05ACE + 0.01ADE - 402.2BCD + 139.9BCE + 27.72BDE + 26CDE + 0.56ABCD - 0.173ABCE - 0.03ABDE - 0.03ACDE - 145.23BCDE + 0.16ABCDE$$
(6)

$$Y_{2} = 31.3 - 0.01A - 37.7B - 43.37C - 10.87D + 1.3E + 0.03AB + 0.04AC + 0.01AD + 0.003AE + 128.3BC + 30.1528BD - 11.5BE + 95.75CD - 10.6CE - 5.4DE - 0.1ABC - 0.03ABD + 0.01ABE - 0.13ACD + 0.02ACE + 0.007ADE - 272.9BCD + 74.17BCE + 23.01BDE + 12.8CDE + 0.3ABCD - 0.11ABCE - 0.04ABDE - 0.02ACDE - 83.82BCDE + 0.12ABCDE$$

$$(7)$$



**Figure 6.** Parity charts of predicted versus observed responses for the initial models of (**a**) phenol conversion ( $Y_1$ ) and (**b**) hydrogen yield ( $Y_2$ ), including all main and interaction independent variables.

**Table 3.** Analysis of the regression coefficient of the initial model for phenol conversion ( $Y_1$ ).

Response	Model Terms	Effect	Coefficient	<i>p</i> -Value
	Model Equation (6)	-	8.74014	0.000
	Ā	17.704	-0.00216	0.008
	В	-12.201	-39.6608	0.016
	С	8.478	-74.4782	0.026
	D	-11.43	-19.4077	0.018
	Ε	41.859	7.49008	0.002
	AB	-1.491	0.023912	0.495
	AC	0.672	0.083992	0.727
Phenol conversion (%)	AD	-0.518	0.030147	0.784
	AE	10.552	0.000846	0.021
	BC	3.915	247.554	0.229
	BD	-0.632	65.8527	0.742
	BE	-7.508	-30.9289	0.036
	CD	-3.939	133.44	0.227
	CE	4.986	-30.316	0.049
	DE	-7.779	-6.77565	0.031
	ABC	-2.213	-0.22352	0.373

Response	Model Terms	Effect	Coefficient	<i>p</i> -Value
	ABD	0.635	-0.08751	0.74
	ABE	-0.731	0.03016	0.706
	ACD	1.525	-0.20745	0.488
	ACE	0.499	0.046393	0.792
	ADE	0.042	0.005443	0.982
	BCD	-4.111	-402.183	0.219
	BCE	0.96	139.889	0.632
Phenol conversion (%)	BDE	-0.528	27.7172	0.78
	CDE	-1.716	25.9847	0.451
	ABCD	3.46	0.557974	0.256
	ABCE	-1.642	-0.17284	0.465
	ABDE	0.062	-0.03197	0.973
	ACDE	1.361	-0.02793	0.524
	BCDE	-2.346	-145.229	0.356
	ABCDE	1.47	0.163287	0.501

Table 3. Cont.

 $R^2 = 0.999, R^2(adj) = 0.976.$ 

**Table 4.** Analysis of the regression coefficient of the initial model for hydrogen yield ( $Y_2$ ).

Response	Model Terms	Effect	Coefficient	<i>p</i> -Value
	Model Equation (7)	-	31.3439	0
	Â	8.806	-0.00816	0.009
	В	-6.519	-37.7222	0.027
	С	4.431	-43.3667	0.041
	D	-7.644	-10.8661	0.018
	E	19.644	1.29028	0.006
	AB	-2.494	0.032778	0.268
	AC	-0.669	0.042667	0.657
	AD	-0.669	0.014172	0.657
	AE	5.019	0.003003	0.039
	BC	1.581	128.333	0.392
	BD	-0.294	30.1528	0.837
	BE	-3.581	-11.5278	0.048
	CD	-2.444	95.75	0.273
	CE	2.644	-10.6167	0.255
Hydrogen vield (%)	DE	-5.506	-5.40194	0.031
riyulogen yielu (70)	ABC	-0.919	-0.1	0.562
	ABD	-0.519	-0.03181	0.724
	ABE	-2.606	0.012222	0.258
	ACD	1.006	-0.1295	0.534
	ACE	-0.456	0.022333	0.753
	ADE	-0.431	0.007131	0.766
	BCD	-0.994	-272.917	0.538
	BCE	0.269	74.1667	0.85
	BDE	-0.181	23.0069	0.898
	CDE	-1.406	12.7861	0.428
	ABCD	2.456	0.345833	0.272
	ABCE	-0.831	-0.10833	0.593
	ABDE	-0.906	-0.03493	0.567
	ACDE	0.894	-0.02364	0.571
	BCDE	-0.181	-83.8194	0.898
	ABCDE	1.119	0.124306	0.497

 $R^2 = 0.998, R^2(adj) = 0.944.$ 

#### 3.3. Reduced Model Equation for Phenol Conversion $(Y_1)$ and Hydrogen Yield $(Y_2)$

In the reduced regression model, all of the insignificant terms were removed and only the variables with a *p*-value of lower than 0.05 were considered. The Pareto chart of the model after reduction at p = 0.05 with 10 degrees of freedom (DF) for  $Y_1$  and nine for  $Y_2$  is presented in Figure 7. All the values were greater than 2.07 (p = 0.05). Moreover, the model is statistically significant as indicated by a *p*-value less than 0.05 (95% confidence).

The reduced linear regression model for production of hydrogen from phenol was fitted for  $Y_1$  and  $Y_2$  through the application of multiple regression analysis on the experimental data and subsequent removal of insignificant terms as shown in Equations (8) and (9). From the analysis of the regression coefficient, the value of the model helps in making inferences about whether there is a need for a complex model for better fit. The significances of the models were determined by the experimental *F*-value derived from the regression coefficient analysis. On the other hand, the tabulated *F*-values are calculated based on the degree of freedom for regression and residual at a specific confidence level. If the *F* test for the model is significant at the 5% level (p < 0.05), then the model is fit and can adequately account for the variation observed.



**Figure 7.** Pareto chart of the reduced model for (**a**)  $Y_1$  and (**b**)  $Y_2$ .

The ratio of the mean square due to regression to that due to error is depicted by the *F* value. Generally, a good model should have its calculated *F* value several times greater than the tabulated. Tables 5 and 6 depict the reduced model coefficients of the factors and interactions, and the *p*-values of the effects in the model are shown for  $Y_1$  and  $Y_2$ , respectively. According to the analysis of the regression coefficient of the reduced regression models, the models are highly significant. This was evident from the Fisher (*F*) test. The ratio of the mean square regression to mean square residual was 121.1 for  $Y_1$  and 60.13 for  $Y_2$  responses. Also, the probability value was very low (*p*-value = 0.000).

$$Y_1 = -1.87 + 0.01A - 3.13B + 4.99C + 0.24D + 1.18E + 0.01AE - 6.3BE + 6.23CE - 1.94D$$
(8)

$$Y_2 = 15.7 + 0.004A - 3.8B + 22.16C + 0.62D + 1.35E + 0.004AE - 2.98BE - 1.38CE$$
(9)

The coefficient of determination ( $R^2$ ) and adjusted  $R^2$  are among other criteria that are commonly used to show the best fitness of a model. Tables 5 and 6 show the calculated  $R^2$  and adjusted  $R^2$  values for the reduced models developed with respect to phenol conversion ( $Y_1$ ) and the hydrogen yield ( $Y_2$ ) responses, respectively. From these values, it was clear that the regression model was adequate.

Response	Model Terms	Effect	Coefficient	F-Ratio	<i>p</i> -Value
	Model Equation (8)	-	-1.87079	121.1	0.000
	Â	17.704	0.0062524	78.52	0.000
	В	-12.201	-3.13	37.29	0.000
	С	8.478	4.9969	18.01	0.000
Phonol conversion $(0/)$	D	-11.43	0.23866	32.73	0.000
Phenoi conversion (%)	Ε	41.859	1.18218	438.95	0.000
	AE	10.552	0.00879326	27.89	0.000
	BE	-7.508	-6.25647	14.12	0.001
	CE	4.986	6.23214	6.23	0.021
	DE	-7.779	-1.94471	15.16	0.001

**Table 5.** The regression coefficient value of the reduced model for  $Y_1$ .

 $R^2 = 0.968, R^2(adj) = 0.955.$ 

**Table 6.** The regression coefficient value of the reduced model for  $Y_2$ .

Response	Model terms	Effect	Coefficient	F-Ratio	<i>p</i> -Value
	Model Equation (9)	-	15.7346	60.13	0.000
	Â	8.806	0.00426	39.93	0.000
	В	-6.519	-3.82292	21.88	0.000
	С	4.431	22.1563	10.11	0.004
Hydrogen yield (%)	D	-7.644	0.61562	30.08	0.000
	Ε	19.644	1.35042	198.67	0.000
	AE	5.019	0.004182	12.97	0.002
	BE	-3.581	-2.98437	6.6	0.017
	CE	-5.506	-1.37656	15.61	0.001

The model reduction has little influence on the predicted values of  $Y_1$  and  $Y_2$ , as indicated by the small change of the  $R^2$  values from the analysis of the regression coefficient results (Tables 5 and 6). The validity of the regression model developed was further tested by drawing the parity charts. The parity plot helps in detecting any outliers and provides a prompt indication of the accuracy of the correlation. The observed values and predicted values of the responses are scattered close to the 95% prediction limits with 45°, indicating an almost perfect fit of the developed reduced model. Figure 8 shows the parity plot of the experimental and predicted values for the reduced model. In all the cases, one or two data points out of 32 fall outside the 95% prediction limits (as shown with blue dotted lines).



**Figure 8.** Parity charts of predicted versus observed responses for the reduced models of phenol conversion ( $Y_1$ ) (**a**) and hydrogen yield ( $Y_2$ ) (**b**) with a 95% prediction band.

## 3.4. Effects of Operating Conditions on Reaction Responses

## 3.4.1. Main Effect Plots for $Y_1$ and $Y_2$

Graphs of the main effects were plotted in order to identify the factor level providing the minimum response value, as shown in Figure 9. Figure 9 presents the mean changes that occurred in both responses ( $Y_1$  and  $Y_2$ ) when the levels of the main factors were changed from the lower level, passing through the central point, to a higher level. The reference line represents the overall mean. A main effect is present when the mean response changes across the levels of a factor. This plot shows that the  $Y_1$  and  $Y_2$  changes depend on the levels of factor E (phenol concentration) with the steepest line. The steeper the slope of the line, the greater the magnitude of the main effect. Therefore, the concentration of phenol should be studied in detail. It can be seen that for the temperature there were great changes in the mean phenol conversion  $(Y_1)$  and hydrogen yield  $(Y_2)$ , characterized by a great degree of departure from the horizontal (but less than the phenol concentration main effect). Therefore, a higher temperature is desirable to get higher phenol conversion and hydrogen yield. This is because of the favorable hydrocarbon conversion at high temperatures [56]. Feed flow rate and Ni/Co ratio showed slight negative changes in mean  $Y_1$  and  $Y_2$ .  $Y_1$  and  $Y_2$  decreased from 39.8% and 40.1% to 27.6% and 33.5% with the increase of feed flow rate from 0.16 to 0.46 mL/min, respectively. On the other hand,  $Y_1$  and  $Y_2$  also decreased from 39.5% and 40.6% to 28.02% and 32.97% with the increase of the Ni to Co ratio from 0 to 1. Nianjun et al. [52] used  $Ni_xCo_y/Al_2O_3$  catalysts to investigate the activity of glycerin liquid reforming. They noticed that an increase in Co content results in maximum H<sub>2</sub> and minimum CH<sub>4</sub> production. They mentioned that low content of Ni is unfavorable for carbon–carbon breaking. Consequently, the performance of the reforming process was lowered. Hence, it can be concluded that temperature and phenol concentration variables highly influenced  $Y_1$  and  $Y_2$ , whereas feed flow rate, Ni/Co ratio, and catalyst weight had minor effects due to their lines being less sharp against the *x*-axis. These patterns were previously identified by statistical significance (Pareto charts and analysis of the regression coefficient data).



**Figure 9.** Main effect plot for (**a**)  $Y_1$  and (**b**)  $Y_2$ .

## 3.4.2. Interaction Plots for $Y_1$ and $Y_2$

To visualize the impact of each factor combination and identify which factors were most influential, interaction plots were constructed. These also consider interactions among variables and can be used to optimize the operating parameters in multivariable systems [57]. The mean response of the two factors is depicted by the plot for all the possible combinations of their settings. When there is interaction between the two factors, the lines appear as non-parallel [58]. Figure 10 presents the full interactions effect plot matrix for five factors, temperature (*A*), liquid feed rate (*B*), catalyst weight (*C*), Ni to Co ratio (*D*), and phenol concentration (*E*) at two collocation levels on phenol conversion ( $Y_1$ ) and hydrogen

yield ( $Y_2$ ). An interaction exists between variables when the effect of one variable on the response depends on the level of another variable [59]. Also, the presence of interaction results in the combined effect being more or less than that expected when the effects are individually added together [60]. Figure 10 shows that phenol concentration factor (*E*) is a predominant factor in  $Y_1$  and  $Y_2$ . This means that E interacted strongly with all other factors. There was a strong two-way interaction between the main effects of *A* and *E*, *B* and *E*, *C* and *E*, *D* and *E*, and *C* and *E*, as evident from the non-parallel nature of the effect lines for the interaction on  $Y_1$  and  $Y_2$  responses. However, it can be proved by regression coefficient analysis. The graph also shows that there is a possible interaction between *C* and *B* factors with respect to  $Y_1$  and  $Y_2$  due to their un-parallel lines. The rest of the graphs are parallel, suggesting a general lack of interaction or weak interaction between the different variables with respect to their influence on  $Y_1$  and  $Y_2$ .



**Figure 10.** Full interaction plot matrix for (**a**)  $Y_1$  and (**b**)  $Y_2$ .

## 3.4.3. Surface and Contour Plots for Phenol Conversion ( $Y_1$ ) Response

As is seen from the analysis of the regression coefficient results in Table 5, all of the variables (temperature (A), feed flow rate (B), catalyst weight (C), Ni-Co ratio (D), and phenol concentration (E)) and their interactions (AE, BE, CE, and DE) are statistically significant and affect the phenol conversion  $(Y_1)$  response. The response surface plots are useful in understanding the relationship between the main and interaction effects of the variables [61]. The contour plots are curved lines because the model contains the interactions of the factors (AE, BE, CE, and DE). The surface and contour plots given in Figures 11 and 12 show the relative effects of any two variables when the remaining variables are kept constant. In general, the results of surface and contour plots reveal that the phenol concentration plays an important role in catalyst activity towards phenol conversion. The surface and contour plots for AE in Figures 11a and 12a show the interactive effect of phenol concentration (E) and temperature (A) by keeping B, C, and D constant. The phenol conversion is relatively high for high levels of A and E. In the same way, the surface and contour plot show the interactive effect of B and E through keeping A, C, and D constant. Figures 11b and 12b show the  $Y_1$  response of the mutual interactions between the variables. It was found to be high at a low flow rate and high phenol concentration. Furthermore, when the catalyst weight ranged from 0.1 to 0.3 g the effect on the  $Y_1$  response was significant (Figures 11c and 12c). On the other hand, Figures 11d and 12d show that the NiCo ratio and phenol concentration (D versus E) have a negative effect on  $Y_1$  response. In conclusion, surface and contour plots prove the results of the regression coefficient data and the Pareto charts.



**Figure 11.** Surface plot for significant interactions on  $Y_1$ : (a) *AE*; (b) *BE*; (c) *CE*; (d) *DE*.



**Figure 12.** Contour plot for significant interactions on  $Y_1$ : (a) *AE*; (b) *BE*; (c) *CE*; (d) *DE*.

3.4.4. Surface and Contour Plots for Hydrogen Yield (Y<sub>2</sub>) Response

Figures 13 and 14 give surface and contour plots indicating the effect of interactive effect temperature (*A*), feed flow rate (*B*), catalyst weight (*C*), Ni-Co ratio (*D*), and phenol concentration (*E*) upon the hydrogen yield ( $Y_2$ ) response. As in the previous section, the highest  $Y_2$  values are observed at the highest phenol concentration. Figures 13 and 14 illustrate that E is more significant with a lower level of *B* and *D*, and a higher level of *A* and *C* factors. The effect of two-way interaction on  $Y_2$  response is almost the same as the  $Y_1$  response and proves the analysis of the regression coefficient data in Table 6.



**Figure 13.** Surface plot for significant interactions on *Y*<sub>2</sub>: (**a**) *AE*; (**b**) *BE*; (**c**) *CE*; (**d**) *DE*.



**Figure 14.** Contour plot for significant interactions on *Y*<sub>2</sub>: (**a**) *AE*; (**b**) *BE*; (**c**) *CE*; (**d**) *DE*.

## 3.4.5. Response Optimization

Finally, the Minitab Response Optimizer tool helps to obtain the best possible result for phenol conversion ( $Y_1$ ) and hydrogen yield ( $Y_2$ ). Moreover, it can help solve problems involving several design variables and one or more responses simultaneously. It also has an interactive graphic mode that helps in investigating local behaviors around an optimized solution. By using a response optimizer, the optimal conditions for the suggested reduced model (Equations (8) and (9)) were found to be 800 °C, 0.16 mL/min feed flow rate, 0.3 g catalyst weight, 0 Ni/Co ratio, and 10 wt % of phenol concentration in order to obtain 94.98% of  $Y_1$  and 67.4% of  $Y_2$  responses (Figure 15a,b, respectively).



**Figure 15.** Results from the Response Optimizer tool for  $Y_1$  (**a**) and  $Y_2$  (**b**).

## 3.5. Coke Analysis

The changes occurring during the steam reforming of Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts were studied through thermo-gravimetric analyses (TGA) of the used catalysts for different phenol concentrations. As is shown in Figure 16, they had different amounts corresponding to the weight loss associated with decomposition/oxidation of graphite from the catalyst. The carbon deposit types depend carbon oxidation temperature. The curves of weight loss as a function of temperature can be divided into three parts: below 200 °C (WL<sub>1</sub>), 200–600 °C (WL<sub>2</sub>), and above 600 °C (WL<sub>3</sub>). In the literature, the region below 200  $^{\circ}$ C has been attributed to water and moisture evaporation. The weight loss between 200 °C and 600 °C has been ascribed to less stable deposits [62], whereas the weight loss above 600  $^{\circ}$ C is generally attributed to oxidation of coke deposits with a more stable [62] and/or different degree of graphitization [63]. In general,  $Ni/ZrO_2$  has less coke deposited on its surface than Co/ZrO<sub>2</sub>. The TGA result shows that the Ni/ZrO<sub>2</sub> catalyst for 2% phenol concentration in the feed did not suffer from any coke formation on its surface in all WL<sub>1</sub>, WL<sub>2</sub>, and WL<sub>3</sub> regions. This catalyst profile exhibited additional weight gained; this might be due to the oxidation of the metallic active sites [64,65]. A significant mass loss was found for both Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts when the phenol concentration increased from 2% to 10% and the temperature rose above 600  $^{\circ}$ C due to the increase in carbon molecules in the feed. Therefore, a higher phenol concentration leads to higher catalyst activity in terms of phenol conversion and hydrogen production but causes the carbon deposition to increase.



Figure 16. TGA profile of used Ni/ZrO<sub>2</sub> and Co/ZrO<sub>2</sub> catalysts.

## 4. Conclusions

In this study, hydrogen production from catalytic steam reforming of phenol was performed over Ni/ZrO<sub>2</sub> (Ni-Co ratio = 1) and Co/ZrO<sub>2</sub> (Ni-Co ratio = 0) catalysts. In order to investigate the promising catalyst activity towards phenol conversion ( $Y_1$ ) and hydrogen yield ( $Y_2$ ), as well as reduce the total number of experiments, a statistical design of experiment (DOE) using the 2<sup>5</sup> full factorial method (five variables, each variable at two levels) was performed. The five variables were temperature (A), liquid feed flow rate (B), catalyst weight (C), Ni to Co ratio (D), and phenol concentration (E). Thirty-two sets of experiments were performed to explore the effects of the reaction variables and their interaction on the phenol conversion and the hydrogen yield. The catalyst performance testing shows that 800 °C, 0.16 mL/min feed flow rate, 0.3 g of catalyst, 0 ratio of Ni-Co (Co/ZrO<sub>2</sub>), and 10 wt % of phenol in the feed gave a phenol conversion of 96.8% and a hydrogen yield of 72.4%. The statistical DOE indicated that within the limits that were being examined, all the variables had a significance level of 5% (p < 0.05). In addition, the interaction among A versus E, B versus E, C versus E, and D versus *E* were found to have a significant effect on  $Y_1$  and  $Y_2$  responses. This was proven by Pareto charts and Normal plots. The analysis of the regression coefficient results for the reduced regression model indicates that the model can fit the experimental data for both  $Y_1$  and  $Y_2$  responses with correlation coefficients ( $R^2$ ) of 0.968 and 0.936, respectively. The parity charts of predicted versus observed responses also confirm the fit of the experimental data with the regression model. The p-value of the reduced model was zero and that model is considered to be statistically significant. The surface and contour plots show that the *E* variable is more significant with higher levels of *A* and *C* and lower levels of *B* and *D*. The response optimizer tool indicated that the optimum conditions for the reduced model were 800 °C, 0.16 mL/min, 0.3 g catalyst, 0 Ni-Co ratio, and 10 wt % of phenol concentration to obtain 94.98% of  $Y_1$  and 67.4% of  $Y_2$ , respectively.

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